

Physical activity and cold pain tolerance in the general population

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Significance:

This study finds that higher level of self-reported leisure-time physical activity is associated with increased cold-pressor pain tolerance in a large population-based sample. Though present in both sexes, the association is strongest among men. Despite the robust dose-response relationship between pain tolerance and self-reported activity level, no such relationship was found for accelerometer-measured activity, reflecting a possible discrepancy in the aspect of physical activity measured. Though the study design does not permit causal conclusions, the findings suggest that increasing physical activity may increase pain tolerance in the general population.

Background: The relationship between habitual physical activity (PA) and experimental pain tolerance has been investigated in small samples of young, healthy, and/or single-sex volunteers. We used a large, population-based sample to assess this relationship in men and women with and without chronic pain.

Methods: We used data from the sixth and seventh Tromsø Study surveys (2007-08; 2015-16), with assessed pain tolerance of participants with the cold-pressor test (CPT: dominant hand in circulating cold water at 3°C, maximum test-time 106 seconds), and self-reported total amount of habitual PA in leisure time (n=19,087), exercise frequency (n=19,388), exercise intensity (n=18,393), and exercise duration (n=18,343). A sub-sample had PA measured by accelerometers (n=4,922). We used Cox regression to compare CPT tolerance times between self-reported PA levels. For accelerometer-measured PA, we estimated hazard ratios for average daily activity counts, and for average daily minutes of moderate-to-vigorous PA done in bouts lasting 10 minutes or more. Models were tested for PA-sex, and PA-chronic pain and PA-moderate-to-severe chronic pain interactions.

Results: Leisure-time PA, exercise intensity, and exercise duration were positively associated with CPT tolerance ($p < 0.001$; $p = 0.011$; $p < 0.001$). More PA was associated with higher CPT tolerance. At high levels of leisure-time PA and exercise intensity, men had a significantly higher CPT tolerance than women. Accelerometer-measured PA was not associated with CPT tolerance.

Conclusions: This study is one of the first to show that higher self-reported habitual PA was connected to higher experimental pain tolerance in a population-based sample, especially for men. This was not found for accelerometer-measured PA.

1 Introduction

2 Several reviews summarize how acute bouts of physical activity (PA) reduce sensitivity to
3 experimental pain stimuli, manifested as temporary change in parameters like sensitivity
4 thresholds and tolerance thresholds (Koltyn 2000; Naugle et al., 2012; Rice et al., 2019). This
5 effect, called exercise-induced hypoalgesia, is seen using electrical, heat, cold, chemical, and
6 pressure pain modalities. A recent RCT found reduced pain sensitivity not to depend on
7 intensity of acute exercise alone, but also on underlying fitness status (Schmitt et al., 2020).
8 Indeed, a more enduring pain sensitivity reduction has been suggested as a feature associated
9 with increased levels of *habitual* PA; a long-term counterpart to the transient exercise-induced
10 hypoalgesia. This is seen using a prospective exercise intervention approach (Jones et al.,
11 2014), comparing athletes to non-athletes (Geva and Defrin 2013; Tesarz et al., 2012), or
12 looking at self-reported (Lemming et al., 2015; 2017; Naugle and Riley 2014) or device-
13 measured PA (Ellingson et al., 2012; Naugle et al., 2017; Ohlman et al., 2018), with heat,
14 cold, pressure, or ischemic pain modalities. The hypothesis of a long-term effect of PA on
15 pain sensitivity was also supported by a meta-analysis of observational studies finding lower
16 pain sensitivity in athletes compared to normally active controls (Tesarz et al., 2012).

17 Although an association with acute bouts of PA and even habitual PA seems to be
18 well-founded, studies often examine single-sex samples despite well-established sex-
19 differences in clinical and experimental pain (Mogil 2012; Racine et al., 2012). They are also
20 often based on small, non-generalizable samples of young, healthy volunteers, and
21 infrequently report accelerometer-measured PA.

22 Adverse change in central mechanisms of pain facilitation and inhibition appears to be
23 a recurring component in several chronic pain conditions (Granovsky 2013; Moana-Filho et
24 al., 2018; O'Brien et al., 2018; Yarnitsky 2010), and has accordingly been hypothesized to be
25 an independent risk factor for developing chronic pain (Baert et al., 2016; Petersen et al.,
26 2018; Staud 2012; Treede 2019; Yarnitsky et al., 2008). As habitual PA is an effective
27 treatment modality and has been suggested to prevent chronic pain (Ambrose and Golightly
28 2015; Holth et al., 2008), part of this effect is thought to occur through upregulating pain-
29 inhibiting mechanisms. However, if chronic pain is already present, this might in some cases
30 sensitize individuals to pain in such a way as to act contrary to the benefits of PA on pain
31 sensitivity. Indeed, the presence of chronic pain has been reported to coincide with a lacking,
32 or even reversed, association between habitual PA and pain sensitivity (Mani et al., 2019; Orr
33 et al., 2017), and identical acute exercise regimens can produce different central pain

34 processing responses across different painful conditions (Meeus et al., 2015). It is therefore of
35 interest to further assess how the presence of chronic pain might influence the relationship
36 between levels of habitual PA and the experience of painful stimuli.

37 To improve our understanding of the relationship between habitual PA and pain
38 sensitivity, studies combining heterogeneous study populations with large samples are
39 warranted. The Tromsø Study has accumulated the hitherto largest population-based
40 experimental pain data sample in the world. These data also contain self-reported and
41 accelerometer-measured habitual PA. Thus our objective was to model relationships between
42 types and measurements of PA and experimental pain sensitivity in a population-based
43 sample, including both sexes with and without chronic pain.
44
45

46 **2 Methods**

47 **2.1 Study population and sample**

48 The Tromsø Study, conducted in the Tromsø municipality in Northern Norway, consists of
49 seven repeated surveys from 1974 to 2016 (Tromsø 1-Tromsø 7). It has invited both total
50 birth cohorts and random samples (Eggen et al., 2013; Jacobsen et al., 2012). Participants
51 were recruited through mailed invitations and received no monetary reimbursement for
52 attending. Data have been collected through questionnaires, biological samples, and clinical
53 examinations. Experimental pain testing using the Cold-pressor test (CPT) was included in
54 Tromsø 6 (2007-08) and Tromsø 7 (2015-16). The participation proportion in Tromsø 6 was
55 66% (n=12,984; age 30-87 years, 53% women), and 65% in Tromsø 7 (n=21,083; age 40-99
56 years, 53% women).

57 For this cross-sectional study, we included individuals who participated in CPT in
58 Tromsø 6 or 7 and had provided data on PA (Figure 1). For participants who had provided
59 data in both Tromsø 6 and 7 (n=6,500), we chose to use CPT, exposure, and covariate data
60 from Tromsø 7 only.

61 Second visit: Of all invitees to the first visit of Tromsø 7, a random sample was made
62 of 20% of participants in age groups 40-59 (n=4,008) and 50% of participants in age groups
63 60-84 (n=6,142). In addition, the study invited all other participants of Tromsø 7 who had also
64 participated in select clinical examinations in Tromsø 6 (n=3,154). Of all these invitees to the
65 second visit of Tromsø 7, 63% (n=8,346) participated. The second visit contained more
66 extensive examinations, including measurement of PA by accelerometry (Figure 1).

67

68

Insert Figure 1 approximately here

69

70 **2.2 Measurements**

71

72 **2.2.1 Physical activity**

73 This study used three different methods to assess PA. First, participants self-reported level of
74 leisure-time physical activity (LTPA) using a modified version of the four-category Saltin and
75 Grimby questionnaire (Grimby et al., 2015), which asks for average level of LTPA during the
76 previous 12 months. Respondents can select from 4 mutually exclusive categories: Reading,
77 watching TV, or other sedentary activity; walking, cycling, or other forms of exercise at least
78 four hours a week (with examples); participation in recreational sports, heavy gardening, etc.
79 at least four hours a week; or participation in hard training or sports competitions, regularly
80 several times a week. Second, participants reported habitual exercise frequency (EF – “How
81 often do you exercise”); habitual exercise intensity (EI – “If you exercise – how hard do you
82 exercise”); and habitual exercise duration (ED – “For how long do you exercise (give an
83 average)”). Third, PA was measured by accelerometer in a sub-sample of participants.

84

85 **2.2.1.1 Accelerometer recordings**

86 PA was measured using an ActiGraph wGT3X (ActiGraph Corp, Pensacola, Florida).
87 Participants were asked to wear the accelerometer on the hip for seven consecutive days
88 except during showering/bathing or swimming. Acceleration was measured in three axes at a
89 sampling rate of 100Hz and reduced to counts as a measure of PA. Non-wear time was
90 defined using the Hecht 2009 algorithm (Hecht et al., 2009). According to this algorithm, at
91 least two of the following conditions had to be met for any given minute to classify as valid
92 wear time: 1) >5 counts per minute; 2) at least two minutes with counts>5 in the following 20
93 minutes; 3) at least two minutes with counts >5 in the preceding 20 minutes. For processing
94 of the counts data into variables defining PA levels, we used Quality Control & Analysis Tool
95 (QCAT), a custom-made software developed in Matlab (The MathWorks, Inc., Natick,
96 Massachusetts, USA). For the analyses, two PA variables were used: first, a variable showing
97 the average daily number of accumulated activity counts; second, a variable expressing
98 moderate to very vigorous PA (MVPA) minutes per day occurring in bouts of activity lasting
99 >10 minutes. This categorization of PA intensity was based on a combination of Sasaki et al.

100 and Peterson et al. cut-offs for triaxial counts per minute (Peterson et al., 2015; Sasaki et al.,
101 2011): sedentary <150; light 150-2689; moderate 2690-6166; vigorous 6167-9642; very
102 vigorous >9642. Counts per minute >2690 were aggregated into moderate to very vigorous
103 PA (MVPA).

104 Exclusion criteria from accelerometry were cognitive or physical impairments
105 preventing participants from handling small devices. A total of 6,333 invited individuals
106 consented to participate in accelerometry. We excluded 43 participants due to lost
107 accelerometers and technical errors, 165 participants due to less than four days with at least
108 10 hours of wear time, and 340 participants due to missing CPT data. Thus, the final sub-
109 sample with valid accelerometry included 5,785 individuals (Figure 1). Accelerometer data
110 gathering and variable generation in the Tromsø Study has been extensively described
111 elsewhere (Sagelv et al., 2019).

112

113 **2.2.2 Cold-pressor test tolerance**

114 The outcome of interest, pain tolerance threshold, was measured on-site as tolerance time
115 during the CPT. Participants were asked to place their dominant hand and wrist in a 13-litres
116 plexi-glass vat containing continuously circulated 3.0°C water. Temperature control was
117 provided by an attached cooling circulator (Julabo FP40HE, Julabo Labortechnik GmbH
118 Germany, 22 liters/min) and temperature in the external plexiglass chamber was calibrated
119 with a precision thermometer. Participants were asked to keep their hand open and relaxed
120 and hold it in the water for as long as possible, up to a maximum tolerance time of 106
121 seconds in Tromsø 6 and 120 seconds in Tromsø 7. Since maximum times differed for the two
122 surveys, Tromsø Study tolerance times were censored at 106 seconds post hoc. Participants
123 were informed of the possibility to abort the test at any time should the pain become
124 unbearable. Reasons for exclusion from CPT included participant reluctance; bilateral loss of
125 sensitivity in the hand; conditions causing a breach of the skin (open sores, painful eczema
126 etc.) affecting both hands; Reynaud's syndrome or cold allergy where the participant believed
127 this to be an obstacle for participation, and; inability to comprehend instructions. In instances
128 where individuals were only able to participate with their non-dominant hand, this was
129 allowed. At the CPT station at Tromsø 6, 1,831 participants were not seen due to capacity
130 limitations of the station; in such cases, staff were requested to prioritize participants <60
131 years of age as that was the age-group least sampled in the study (Stabell et al., 2013).
132 Individuals not seen at the station were counted as not having participated in CPT (Figure 1).

133

134 **2.2.3 Covariates**

135 Several covariates were assessed as possible confounders as described below. These were
136 investigated based on a rationale that other works have found such factors to be associated
137 with painful conditions, pain sensitivity, or associated morbidity. We had questionnaire-data
138 on the following covariates: a) education level (primary/secondary school up to 10 years,
139 upper secondary up to three years, college/university less than four years, college/university
140 for four years or more); b) daily smoking (never, former, or current daily smoker) and
141 reporting of number of cigarettes smoked per day for present or former daily smokers,
142 combined in a categorical variable (never smoked daily, smoked daily previously, smokes
143 between one and ten cigarettes daily, smokes more than ten cigarettes daily); c) self-reported
144 health (very bad, bad, neither good or bad, good, excellent), combining “very bad and bad”;
145 and d) alcohol consumption frequency (never, monthly or less, 2-4 times a month, 2-3 times a
146 week, 4 or more times a week), combined with habitual number of units consumed when
147 drinking alcohol (1-2, 3-4, 5-6, 7-9, 10 or more). The information about alcohol consumption
148 frequency and units consumed was used to create a categorical variable of approximate
149 tertiles indicating the average number of units consumed each week. Furthermore, we used
150 waist-height-ratio (WHtR) as an alternative to body-mass index (BMI), calculated by dividing
151 in situ-measured waist circumference in centimeters on body height in centimeters in
152 accordance to Swainson et al. (Swainson et al., 2017).

153 Information on chronic pain was obtained from a yes/no question: “Do you have
154 persistent or constantly recurring pain that has lasted for three months or more”. In Tromsø 7,
155 96% (N=20,263) of participants reported on the absence/presence of chronic pain, as well as
156 distribution and characteristics of all present pain, on an electronic body map, the Graphical
157 index of pain (GRIP) (Steingrímssdóttir 2020). Characteristics included pain location, onset,
158 intensity, impact on activities of daily living, and bothering, for each painful area.
159 Characteristic items included a ‘not applicable’ option for those that had no chronic pain. Due
160 to not participating in Tromsø 7, 2,987 participants of the present study sample had no GRIP-
161 data. For those participating, a technical error during a brief interval of the study period
162 caused the loss of GRIP-data for 642 of the participants in our sample.

163

164 **2.3 Statistical methods**

165 Participant characteristics were described using means and standard deviations (SD) for
166 continuous variables, and proportions for categorical variables. The distribution of CPT

167 tolerance times was right-censored at a value corresponding to the upper time limit for the
168 test. Additionally, 10-minute bout MVPA was right-skewed. We therefore used median and
169 inter-quartile range (IQR) to describe these data.

170 We assessed the association between PA and CPT tolerance using Cox proportional
171 hazard regression models. This is a time-to-event model which estimates group differences in
172 risk of experiencing an adverse event (in our case, the event of withdrawing the hand from the
173 cold water prior to the maximum test-time possible) at any given time during the test. Our
174 group comparison was level of PA. Participants reaching the maximum test-time of 106
175 seconds were right-censored, i.e. they were counted by the model as having been at risk of but
176 not having experienced the event of interest during the test time. As such, the model considers
177 both the number of participants at risk of the event in each group at any given time of CPT, as
178 well as the rates at which participants of each group are experiencing the adverse event during
179 the test. The resulting “hazard rates” of the groups can be compared across groups as “hazard
180 rate ratios” (HRs) which here serve as comparisons of how well participants in different PA
181 groups tolerate the test stimulus. Thus, the HRs are the effect estimates of interest.

182 We used the Schoenfeld residuals test as well as visual inspection of log-log survival
183 plots to ensure that the proportional hazards assumption was not violated – that is, that HRs
184 were not dependent on the time of CPT.

185 Separate models were estimated for each PA exposure (Figure 1). Four models used
186 questionnaire-derived PA as exposure. When estimating models for self-reported PA, we first
187 included exposures as continuous variables to estimate significance of trend. Followingly, the
188 lowest exposure categories were used as reference groups for group comparisons. For self-
189 reported EF and ED, the lowest two exposure categories were combined into single categories
190 to preserve statistical power. Two models were based on data from accelerometry as the main
191 exposure, constituting sub-group analyses. The first of the accelerometry models was fitted
192 using average amount of activity per valid day as the independent variable of interest, where
193 the activity of a valid day was expressed as the average number of counts per minute per day.
194 The other model was fitted using average daily minutes of MVPA done in bouts lasting 10
195 minutes or more as the independent variable of interest. Both accelerometer variables were
196 included as continuous variables and HRs were reported per standard deviation increase.

197 All six models were adjusted for sex and age. Other listed covariates were assessed as
198 possible confounders. Confounding was regarded as present if adding a covariate to any sex-
199 and age-adjusted model changed the exposure-outcome coefficient by more than 10% in
200 either direction. If confounding was regarded as present in any model, the confounder was

201 included in all models.

202 To assess the impact that chronic pain might have on the PA-pain tolerance
203 association, we tested for the presence of a chronic pain·PA interaction by including a two-
204 way cross product term in our regression models and assessing its statistical significance. We
205 did the same for two-way cross product terms of sex·PA. We then used likelihood ratio tests
206 to compare model fit with and without interaction terms. If interaction with chronic pain was
207 present, models were presented stratified according to chronic pain status.

208 We performed a sensitivity analysis to assess the impact of different definitions of
209 chronic pain when assessing interactions between PA and chronic pain. This was done by
210 comparing a “chronic pain yes/no” question from both Tromsø 6 and 7, to a “moderate-to-
211 severe chronic pain” item. To create this, we used a combination of the Tromsø 7 GRIP pain
212 characteristics as an approximation of the ICD-11 criteria regarding intensity, bothering, and
213 impact of moderate-to-severe chronic pain (Treede et al., 2019): onset \geq 3months, intensity
214 >3 , bothering >3 , impact on ADL >3 (all on a 0-10 numeric rating scale). Some participants
215 had missing information on some of these characteristics (not including participants
216 responding ‘not applicable’). Therefore, we compared the complete cases-model of moderate-
217 to-severe chronic pain to a model which imputed missing GRIP data, as described below.

218 Another sensitivity analysis examined the associations between LTPA and CPT
219 tolerance in the accelerometry sub-sample, to see whether the association differed in the sub-
220 sample compared to the sample of the LTPA model.

221 All HRs are reported with 95% confidence intervals (CIs), and the significance level
222 was set at 5%. Data analyses were performed using STATA 15.0 (StataCorp, College Station,
223 TX, USA).

224

225 **2.4 Missing and multiple imputation**

226 Appendix Table S1 shows frequencies and proportion of missing on covariates. Most of the
227 missing information was attributable to item non-response of PA and chronic pain. To assess
228 the impact of missing data on results, and to include observed data otherwise lost to analysis,
229 we imputed missing covariable data for the models of LTPA, EF, EI, and ED. When
230 compared, results from imputation generally yielded small differences to our complete cases-
231 models. The one notable difference was one level of one exposure for women changing from
232 borderline non-significant to statistically significant (Appendix Table S2). Henceforth, we
233 present results from complete-cases models only. Figure 1 shows number of participants

234 included in complete case model after excluding for all types of missing.

235 We also imputed GRIP-values for those participants who reported pain in the GRIP of
236 Tromsø 7 but were missing information on one or more of the pain characteristics required to
237 compute the moderate-to-severe chronic pain variable. We then compared the model based on
238 imputed values to that of the complete-cases model. Multiple imputation was performed using
239 chained equations on 100 imputed datasets with predictive mean matching (known nearest
240 neighbors=10).

241

242 **2.5 Ethics**

243 The current study was approved by the Regional Ethics Committee of North-Norway (ref.
244 REK North 2016/1794). All participants gave written informed consent. Data from three
245 participants who withdrew their consent were not used in the analysis.

246

247 **3 Results**

248 Baseline characteristics for study participants are given in Table 1. In total 22,271 individuals
249 participating in CPT in either Tromsø 6 or Tromsø 7 were included in the analyses. Of these,
250 12,881 (58%) of participants, of whom 57% were women, withdrew their hand before the
251 maximum test time of 106 seconds. Total median CPT tolerance was 49 seconds for women
252 and 95 seconds for men. Median CPT tolerance for only those participants who withdrew
253 their hand was 32 seconds (IQR 27); 30 seconds for women (IQR 27), and 34 seconds for
254 men (IQR 28).

255 According to accelerometry-measured PA, median daily amount of MVPA performed
256 in bouts of 10 minutes or more was 7.6 minutes (IQR 19.7). Table 1 further shows mean valid
257 wear-days and wear-time in hours per day. The sub-group with accelerometry measurements
258 was on average six years older than the main study sample.

259

260 *****Insert Table 1 approximately here*****

261

262 **3.1 Self-reported PA and CPT tolerance**

263 Figure 2 shows the proportion of participants who aborted CPT before the maximum time or
264 who were right-censored, by LTPA level at intervals of CPT tolerance time. Compared to the

265 sedentary participants, all higher LTPA categories were significantly associated with higher
266 CPT tolerance (Table 2). We observed a significant interaction between PA and sex, with an
267 additional increase in pain tolerance with higher PA level for males. Only women who
268 reported vigorous LTPA showed a significant increase in CPT tolerance compared to women
269 reporting sedentary LTPA. In sex-specific analyses, associations were stronger with larger
270 effects for men than women although, in this one instance, the effect for women was larger
271 than for men. Table 2 further shows that EF for both sexes combined was not significantly
272 associated with CPT tolerance at any level of exposure, although the direction of the effect
273 was consistent with that of other exposures. Moderate EI was significantly associated with
274 higher CPT tolerance compared to light EI. Analysis showed a significant interaction between
275 moderate EI and sex, and sex-specific analysis revealed that the association was significant
276 for males only. The highest two levels of ED were significantly associated with higher CPT
277 tolerance compared to the level of shortest duration. Analysis showed no significant
278 interaction between ED and sex, and results were significant for both sexes when analysed
279 separately.

280 All significant HRs were smaller than 1, with all directions of effect indicating
281 increased CPT tolerance with higher PA.

282

283 **3.1.1 Chronic pain and CPT tolerance**

284 Of the 18,642 participants of CPT that responded to GRIP, a total of 2,022 participants had
285 missing data on either time of onset, intensity, bothering, or impact on activities of daily
286 living for any area they reported to be painful. This left 16,620 participants with complete
287 GRIP information on chronic pain prevalence as well as chronic pain characteristics,
288 including those responding ‘not applicable’, from which to construct the moderate-to-severe
289 chronic pain item (Table 1). Using this definition of chronic pain, the prevalence of chronic
290 pain among the respondents of GRIP was 18,4%.

291 Results from two-way interaction analyses between PA and chronic pain on CPT
292 tolerance are presented in table S3, and between PA and moderate-to-severe chronic pain on
293 CPT tolerance in table S4.

294 We found indication of an interaction with chronic pain on the relationship
295 between EI and CPT tolerance. This was found using both the simple item no chronic pain
296 versus chronic pain (pain duration \geq 3months), and moderate-to-severe chronic pain as
297 defined according to the criteria suggested in ICD-11. Specifically, we found significant

298 interaction effects for those who exercised at vigorous intensity. In individuals with chronic
299 pain we observed a stronger, positive association between EI and pain tolerance compared to
300 those reporting no chronic pain. Despite no significant complete-case interactions between
301 ED and moderate-to-severe chronic pain, the imputed model found a significantly stronger
302 association with CPT tolerance for the highest level of ED for those without pain (Table S4).

303

304 **3.2 Accelerometer-measured PA and CPT tolerance**

305 HRs for total counts and 10-minute bout MVPA minutes are reported in Table 2. Associations
306 between accelerometer-measured PA and CPT tolerance were not statistically significant. We
307 found no interaction with sex or chronic pain.

308 Differences in associations of self-reported LTPA and CPT tolerance between the
309 main sample and the sub-group with accelerometry data were found to be negligible (results
310 not shown).

311

312 *****Insert Table 2 & Figure 2 approximately here*****

313

314

315

316 **4 Discussion**

317 In this study, self-reported LTPA, EI, and ED were positively associated with CPT tolerance
318 in a dose-response relationship whilst accelerometer-measured PA was not. Chronic or
319 moderate-to-severe chronic pain did not moderate these relationships, suggesting the
320 association between PA and pain tolerance to remain independent of either in this sample.

321

322 **4.1 PA and pain tolerance**

323 Reviews have summarized possible mechanisms through which acute PA might affect pain
324 sensitivity (Rice et al., 2019; Sluka et al., 2018), including activation of endogenous opioid or
325 non-opioid pain-inhibitory systems influencing central mechanisms of pain modulation,
326 regulation of inflammatory mediators, and autonomic nervous regulation of stress response
327 systems. Others have further suggested cardiovascular interactions (Koltyn and Umeda 2006;
328 Ring et al., 2008). These mechanisms may plausibly be involved in long-term effects of PA
329 on pain sensitivity, alongside select psychological factors that may beneficially modulate pain
330 (Baker and Kirsch 1991; Geva and Defrin 2013; Jones et al., 2014). Regardless, the effect of

331 long-term PA on pain sensitivity is surely multifaceted.

332 Previous studies suggest a link between habitual PA and experimental pain tolerance,
333 both when comparing athletes to non-athletes (Geva and Defrin 2013; Tesarz et al., 2012),
334 when comparing self-reported PA levels (Lemming et al., 2015; 2017; Naugle and Riley
335 2014), or measuring PA using accelerometry (Ellingson et al., 2012; Naugle et al., 2017;
336 Ohlman et al., 2018). Jones et al. found increased pain tolerance in a controlled trial following
337 a six-week program of structured moderate to vigorous aerobic cycling (Jones et al., 2014),
338 indicating that change in exercise at a certain level positively influences pain tolerance.
339 Indeed, underlying level of physical fitness is found to affect pain sensitivity independently of
340 acute exercise intensity (Schmitt et al., 2020), although most consistently when looking at
341 pain tolerance thresholds (Tesarz et al., 2012). Schmitt et al. suggested that this reflects a
342 functional adaptation of central neurological mechanisms, explaining why PA is a possible
343 therapeutic avenue towards prevention and regulation of chronic pain conditions.

344

345 **4.1.1 Accelerometer-measured and self-reported PA**

346 In addition to varying according to pain sensitivity parameter studied, correlations between
347 PA and pain sensitivity vary considerably when PA is accelerometer-measured (Black et al.,
348 2017; Ellingson et al., 2012; Ohlman et al., 2018; Waller et al., 2019). One large-sample study
349 found negative, and a lack of, associations between higher levels of accelerometer-measured
350 PA and pain thresholds among 22 year-olds (Waller et al., 2019). Comparing participants with
351 varying distributions of current pain, they found ambiguous associations with pressure and
352 cold pain threshold when measuring PA using an Actigraph GT3X in a scheme much
353 resembling that of our study. Others found significant prediction of pressure-pain threshold by
354 accelerometer-measured MVPA, but no such effect for heat pain threshold (Ohlman et al.,
355 2018).

356 Accelerometry is a feasible large-scale alternative to energy expenditure estimation
357 using more expensive gold-standard measures (Sylvia et al., 2014). Validating triaxial
358 ActiGraph PA intensity cut points against indirect calorimetry, Santos-Lozano et al. found a
359 moderate to high ability to correctly classify PA intensities (Santos-Lozano et al., 2013).
360 Nevertheless, accelerometry might underestimate volume of certain types of PA and their
361 intensity, especially in free-living. For example, the uniaxial ActiGraph MTI seems prone to
362 misclassification of activities such as carrying heavy loads, swimming, or riding a bike,
363 causing underestimation of total energy expenditure (Hagstromer et al., 2007). Also,

364 accelerometer data rarely distinguish between occupational PA and LTPA. Although we are
365 unaware of studies investigating associations between occupational PA and pain tolerance,
366 several have suggested high occupational PA as a risk factor for clinical pain (Bergmann et
367 al., 2017; Heuch et al., 2017; Miranda et al., 2008; Shieh et al., 2016; Sim et al., 2006). Given
368 a link between clinical and experimental pain, this could weaken associations in our study as a
369 possibly detrimental effect of occupational PA counterbalances the effect of LTPA. Finally,
370 there remains variability in accelerometer types, what output they provide, and their
371 corresponding validity in detecting PA correctly (Plasqui et al., 2013).

372 There is also a known discrepancy between self-reported and accelerometer-measured
373 amount of PA in general (Skender et al., 2016) and in the Tromsø Study in particular (Sagelv
374 et al., 2020). Known challenges to questionnaire reliability, validity, and sensitivity include
375 longer periods of recall, low sensitivity to change in patterns of activity or activity-related
376 differences in health, and large errors of absolute estimates of amount of activity (Lee et al.,
377 2011; Shephard 2003; Sylvia et al., 2014), with indications of significant overestimation of
378 volume of PA, in particular higher intensities, with self-report compared to accelerometry
379 (Dyrstad et al., 2014; Hagstromer et al., 2007). Our main analyses ranked and compared
380 activity levels based on self-reported PA. Sagelv et al. found that associations between self-
381 reported PA ranks and accelerometry measures were consistently and significantly positive,
382 although correlations with accelerometer-measured steps, types of PA intensity counts, and
383 bouts of MVPA were negligible to moderate. The Saltin-Grimby PA levels scale correlates
384 well with both VO₂ max, resting heart rate (Emaus et al., 2010), and physical fitness as work
385 capacity (Lochen and Rasmussen 1992), and is significantly associated with risk of
386 myocardial infarction and death (Calais et al., 2014). Although volume of PA can be
387 overestimated, the scale shows high predictive validity, with PA levels consistently inversely
388 associated to “different risk factors, morbidity and health as well as future mortality” (Grimby
389 et al., 2015). While accelerometers seem suitable for measuring PA time·intensity,
390 questionnaires appear useful in ranking and comparing participants’ relative activity levels. In
391 our self-report models we observed a dose-response relationship of long-term PA rank and
392 pain tolerance.

393
394 Utilizing accelerometer-measured PA, our sub-group analysis did not support findings from
395 self-reported PA, despite similar associations of self-reported LTPA and CPT tolerance in the
396 primary sample and sub-groups. The cause of this discrepancy is unknown. It might reflect
397 the difference inherent in assessing energy expenditure and fitness versus ranking PA habits

398 and lifestyles. Although self-report results showed associations between habitual PA and pain
399 tolerance, we cannot accurately state the inherent PA volume and intensity, and whether there
400 is some other quality to an active lifestyle in our participants that mediates this association.
401 No current measurement tool captures all components inherent to PA: intensity, duration,
402 frequency, volume, domain, and context (Sagelv et al., 2020). Rather, methodologies differ
403 with regards to strengths and weaknesses. Future studies should be mindful to select
404 measurements suitable to subject-matter requirements, and should also be aware of possible
405 differences between LTPA and occupational PA. Thus, beyond adding towards confirming a
406 relationship between PA and pain tolerance, our study found those reporting to habitually
407 engage in PA with higher intensities and durations to be most tolerant to pain. This indicates a
408 ‘chronic’ equivalent to the finding by Schmitt et al. of a similar response to both acute
409 exercise and underlying fitness (Schmitt et al., 2020).

410

411 **4.1.2 Sex differences**

412 Reviews and later studies find sex differences in experimental pain, with women generally
413 being more pain sensitive (Bartley and Fillingim 2013; Bulls et al., 2015; Defrin et al., 2009;
414 Hashmi and Davis 2014; Lemming et al., 2015; 2017; Mogil 2012). In a review from 2012,
415 80% of studies looking at CPT found lower cold pain tolerance in women than men (Racine et
416 al., 2012). In our study, men had almost twice the median tolerance time of women, with
417 women more likely to abort the CPT before the maximum test-time. Theories regarding
418 underlying mechanisms of sex-differences in pain have been summarized elsewhere (Bartley
419 and Fillingim 2013; Defrin et al., 2009; Mogil 2012; 2018; Sorge and Totsch 2017), and
420 include sex-dependent differences in immunologic and inflammatory mediation of pain
421 (Mapplebeck et al., 2016; Sorge et al., 2011). In our study, PA was more strongly associated
422 with pain tolerance in men than women. Possible explanations for the sex-specific effect of
423 PA include sex-dependent dimorphism of opioid receptors and descending pain-modulatory
424 circuits (see review (Mogil 2018); (Chakrabarti et al., 2010; Liu and Gintzler 2000; Loyd and
425 Murphy 2014; Tershner et al., 2000)), both of which are mechanisms implicated in the
426 hypoalgesic effect of PA (Koltyn et al., 2014; Naugle et al., 2012; Rice et al., 2019).

427

428 **4.1.2 Chronic pain**

429 Only the level of most vigorous EI had any statistically significant interaction with chronic
430 pain, suggesting even higher pain tolerance when exercising vigorously for those suffering
431 from chronic pain compared to those who were pain-free. In general, we found that dose-
432 response relationships between self-reported PA and pain sensitivity remained with and
433 without chronic or moderate-to-severe chronic pain. Vaegter et al. found increased pain
434 tolerance after acute exercise in subjects with and without, but other experimental pain
435 measures were dependent on the underlying pain sensitivity of patients (Vaegter et al., 2016).
436 Other studies have found inconsistent associations between exercise or self-reported PA and
437 temporal summation of pain or conditioned pain modulation in chronic pain patients (Mani et
438 al., 2019; Meeus et al., 2015; Orr et al., 2017). Similar to the findings of Vaegter et al.
439 regarding acute exercise, our study found a positive relationship between habitual exercise
440 and pain tolerance in pain-free subjects and subjects reporting various forms of chronic pain.
441 The lack of moderating effect by chronic pain on the relationship between PA and pain
442 tolerance indicates that this relationship remains the same for chronic pain-sufferers as for the
443 pain-free, suggesting that PA might still be able to positively influence habitual central
444 modulation of pain despite the presence of chronic pain. However, the present study looks at
445 two dichotomized types of chronic pain in sub-groups that are possibly quite heterogenous.
446 As the association between PA and clinical pain can differ between different types and
447 severities of chronic pain conditions, we might therefore not be able to detect moderation at a
448 more clinically meaningful level. To amend this, future population studies could group results
449 on specific clinical pain states or could stratify analyses according to chronic pain
450 characteristics such as distribution of painful sites. Finally, the link between experimental
451 pain and clinical pain remains to be clarified. Future studies need to assess whether and to
452 what extent pain sensitivity mediates a positive effect of PA on clinical pain states.

453

454 **4.2 Strengths and limitations**

455 The main strength of this study is its unprecedented sample, enabling analysis of habitual PA
456 and pain tolerance in a population-based sample of women and men, with a high participation
457 proportion and with a heterogenous combination of demography and health states, allowing a
458 robust adjustment for possible confounders.

459 Analyses contained both self-reported and accelerometer-measured PA, both of which
460 are methods with known methodological challenges. In addition, accelerometry was not able
461 to distinguish between occupational and leisure-time PA. Another limitation is scarce

462 evidence regarding the reliability of the CPT tolerance parameter. Looking at intra-class
463 correlation coefficients for CPT duration (i.e. tolerance time), one reliability study including
464 19 pain-free students found fair coefficients for test-retest reliability and poor to excellent
465 coefficients for inter-examiner reliability (O'Neill and O'Neill 2015). Koenig et al. reported an
466 intraclass correlation of 0.92 for pain tolerance measured with 4°C CPT at two occasions
467 separated by two weeks in, predominantly female, students (Koenig et al., 2014). Finally, our
468 measure of chronic or moderate-to-severe chronic pain was of low resolution, possibly
469 leading to a heterogenous chronic pain sub-sample and diluted effects of the moderation
470 analyses.

471

472 **4.3 Conclusion**

473 In this population-based study, higher self-reported habitual PA was associated with higher
474 experimental pain tolerance. This association was more evident for men than for women and
475 was dose-response shaped. There were indications of higher tolerance with vigorous exercise
476 for participants with chronic pain. Future studies could further investigate possible
477 relationships between accelerometer-measured LTPA, as well as occupational PA, and pain
478 tolerance.

479

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484 Study, without whom it would not be possible to do what we do.

485

486 **Conflicts of interest**

487 All authors declare that they have no conflicts of interests related to this study.

488

489 **Author contributions**

490 APÅ, CSN, AS, MKF, LAH, AH, BM, and ÓAS all contributed to the collection of data.
491 APÅ and ÓAS planned and outlined the manuscript. APÅ and TW were responsible for the
492 statistical modelling, and APÅ performed all statistical analyses. All authors have contributed
493 to the interpretation and discussion of results, and to the development of the manuscript
494 through critical revision and comments. All authors have approved this paper.

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719 **Figure Legends:**

720 **Figure 1:** Flow of study participants. CPT: cold-pressor test; LTPA: leisure-time physical
721 activity; EF: exercise frequency; EI: exercise intensity; ED: exercise duration. The Tromsø
722 Study 2007-2016.

723 * 644 participants had missing data on one or more PA questionnaires.

724 **Figure 2:** Proportions aborting cold-pressor test and right-censoring over leisure-time
725 physical activity groups; n=21,355. The Tromsø Study 2007-2016.

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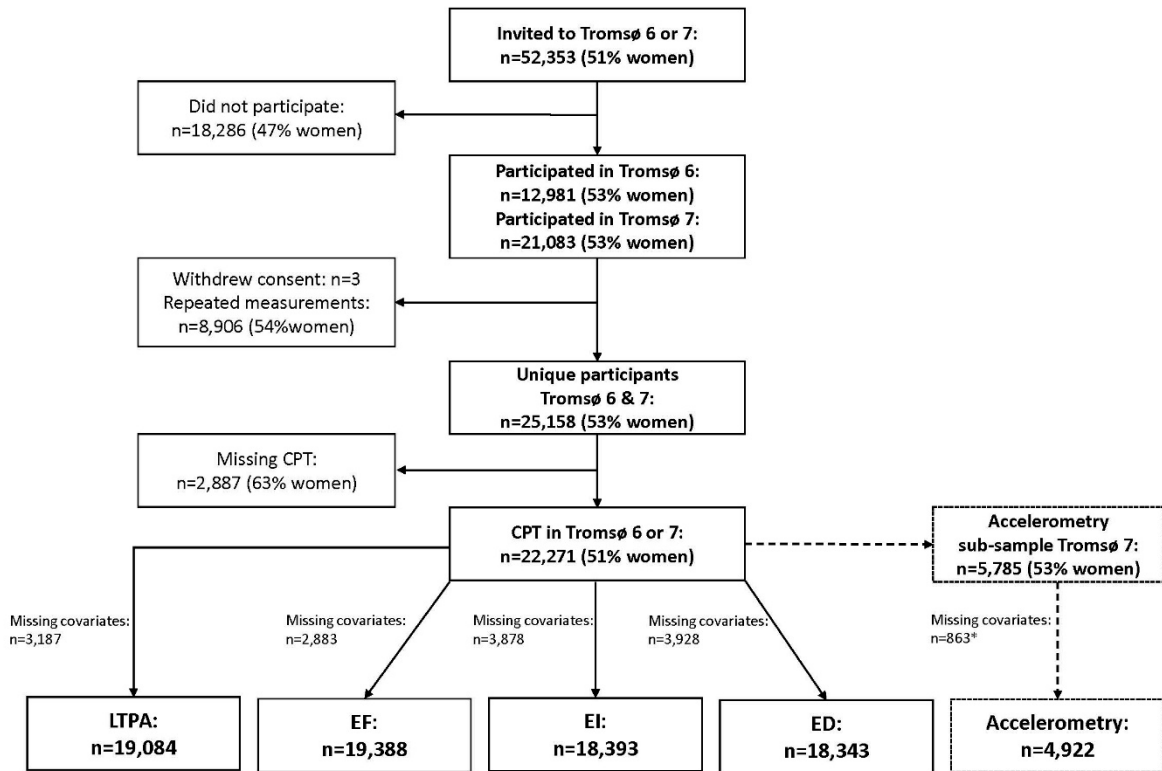
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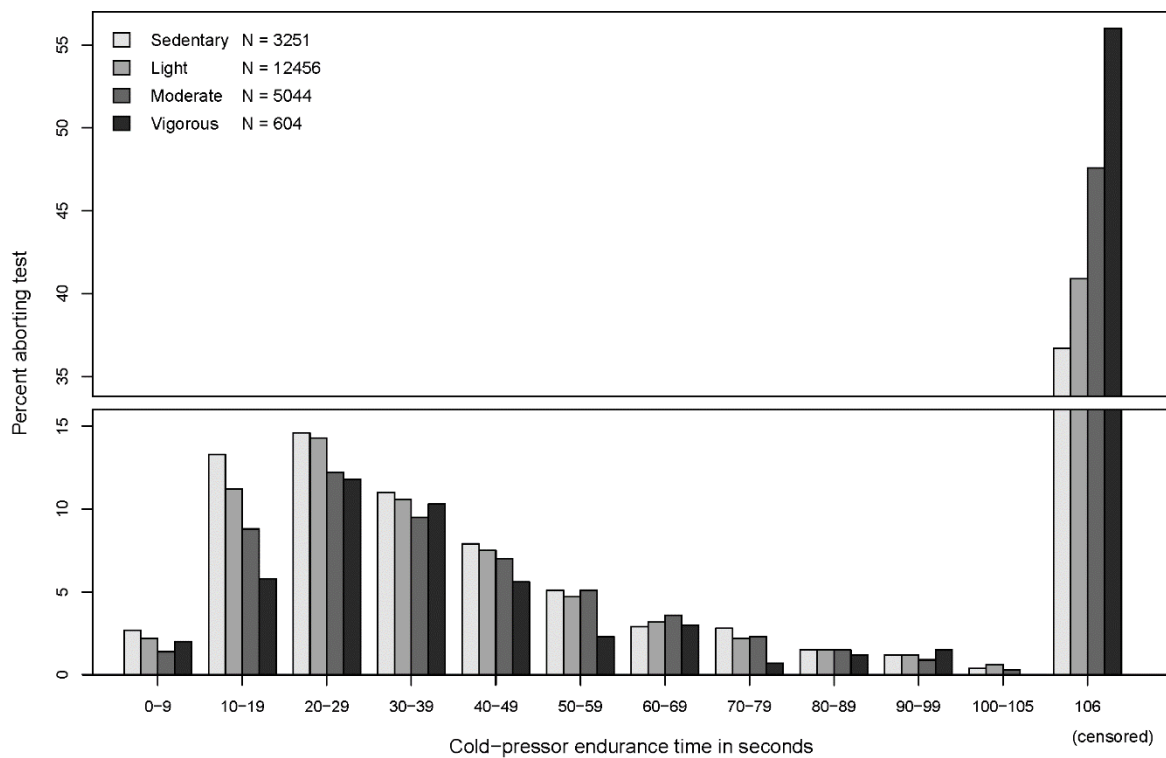


Table 1: Descriptive characteristics of study participants (n=22,271). The Tromsø Study 2007-2016.

Covariate	All	Accelerometry, sub-sample	Withdrew hand in CPT (CPT < 105.6 sec.)	Endured CPT (CPT = 105.6 sec.)
Number of participants (%)	22,271	5,785 (26)	12,881 (58)	9,390 (42)
% Female	51	53	57	43
CPT tolerance time (seconds), median (IQR)	62.5 (76.9)	57.1 (77.8)	31.9 (27.3)	-
Females	49.0 (8.5)	48.7 (80.6)	30.0 (26.9)	-
Males	95.3 (71.8)	71.3 (73.5)	34.3 (27.5)	-
Age, mean (SD)	57.0 (11.6)	63.0 (10.1)	57.0 (11.5)	57.0 (11.8)
WHtR, mean (SD)	0.56 (0.07)	0.56 (0.07)	0.56 (0.07)	0.56 (0.07)
Education level (%):				
Primary/secondary school, up to 10 years	24	28	25	22
Upper secondary, up to 3 years	29	29	30	29
College/university, less than 4 years	19	19	18	20
College/university, 4 years or more	28	24	27	30
Chronic pain (%)	36	35	38	33
GRIP ^a	16,620	5,021	10,001	6,619
GRIP moderate-to-severe chronic pain (% ^b)	3,056 (18.4)	891 (17.8)	2,063 (20.6)	993 (15)
Smoking (%):				
Never	41	39	38	45
Smoked daily previously	44	49	46	41
Smokes 1-10 cigs a day	9.5	8	10	9
Smokes > 10 cigs. a day	5.5	4	6	5
Average alcohol consumption (%):				
Never	8	8	9	8
0.375-0.875 units per week	23	23	24	22
1.125-2.5 units per week	24	25	23	24
>2.625 units per week	46	44	45	47
Self-reported health (%):				
Bad or very bad	5	4	6	4
Neither or	26	27	28	25
Good	54	56	53	55

Excellent	15	13	13	16
Physical activity leisure time (%):				
Sedentary	15	13	17	13
Light	58	62	60	56
Moderate	24	24	21	27
Vigorous	3	2	2	4
Exercise frequency (%):				
Never or less than once per week	17	16	17	16
1-3 times per week	57	56	57	56
Approximately every day	26	28	26	27
Exercise intensity (%) ^c :				
Light	40	44	42	37
Moderate	56	53	54	58
Vigorous	4	3	4	5
Exercise duration (%) ^c :				
0-29 minutes	21	20	22	18
30-60 minutes	57	57	57	57
More than 60 minutes	22	23	21	25
Accelerometry ^d :				
Daily total counts (mean (SD))	-	536 (178)	530 (177)	543 (180)
Daily 10-minute MVPA (median (IQR))	-	7.6 (19.7)	6.9 (18.7)	8.9 (21)
Valid wear-days (mean (SD))	-	6.8 (0.5)	6.8 (0.5)	6.8 (0.5)
Wear-time hours per day (mean (SD))	-	17.3 (1.8)	17.3 (1.8)	17.3 (1.9)

^a Number of non-missing respondents to the Graphical Index of Pain characteristics of time of onset, pain intensity, pain distress, and impact on activities of daily living; includes those without present chronic pain responding 'not applicable' to characteristics.

^b 3,056 / 16,620; 891 / 5,021

^c Habitually, whenever exercising.

^d n=5,785

CPT = Cold-pressor test; IQR = interquartile range; SD = standard deviation; WHtR = waist-to-height ratio; MVPA: Moderate to very vigorous physical activity.

Table 2: Hazard ratios of hand withdrawal on cold-pressor test tolerance according to levels of physical activity by sex^a. The Tromsø Study 2007-2016.

PA type	n =	HR all (CI)	p, trend	HR women (CI)	HR men (CI)
Leisure-time PA, per unit	19,084	0.91 (0.89-0.94)	<0.001	0.93 (0.89-0.97)	0.90 (0.86-0.94)
Sedentary	2,872	1		1	1
Light	11,151	0.91 (0.86-0.96)		0.95 (0.89-1.03)	0.86 (0.79-0.93)
Moderate	4,509	0.85 (0.79-0.90)		0.91 (0.83-1.00)	0.78 (0.71-0.86)
Vigorous	552	0.71 (0.62-0.82)		0.63 (0.51-0.78)	0.81 (0.67-0.97)
Exercise frequency, per unit	19,388	0.98 (0.95-1.01)	0.146	0.96 (0.92-0.997)	1.00 (0.96-1.05)
< 1/week	3,187	1		1	1
1-3 times/week	11,094	0.99 (0.94-1.05)		0.99 (0.92-1.07)	0.99 (0.92-1.07)
Approximately every day	5,107	0.96 (0.90-1.02)		0.93 (0.85-1.01)	1.00 (0.91-1.10)
Exercise intensity, per unit	18,393	0.95 (0.92-0.99)	0.011	0.97 (0.92-1.02)	0.94 (0.89-0.99)
Light	7,212	1		1	1
Moderate	10,402	0.95 (0.91-0.99)		0.96 (0.91-1.02)	0.92 (0.86-0.98)
Vigorous	779	0.94 (0.84-1.04)		0.95 (0.81-1.11)	0.93 (0.81-1.08)
Exercise duration, per unit	18,343	0.91 (0.88-0.93)	<0.001	0.92 (0.89-0.96)	0.89 (0.85-0.93)
0-29 min.	3,681	1		1	1
30-60 min.	10,596	0.86 (0.82-0.90)		0.87 (0.81-0.93)	0.85 (0.78-0.91)
>60 min.	4,066	0.82 (0.77-0.87)		0.85 (0.79-0.93)	0.79 (0.73-0.87)
Accelerometry:	4,922				
Daily total counts^b		0.99 (0.95-1.03)	0.734	1.02 (0.96-1.08)	0.96 (0.91-1.02)
Daily 10-minute MVPA^b		0.98 (0.94-1.02)	0.218	1.00 (0.94-1.05)	0.95 (0.90-1.01)

^a Cox proportional hazards regression.

^b Hazard ratio for 1SD increase

Unstratified models are adjusted for: sex, age, waist-height-ratio, education, current smoker status, average weekly alcohol consumption, self-reported health and chronic pain. Statistically significant results denoted by **bold**. Disregarding sex, stratified models use identical adjustments.

PA: physical activity; MVPA: moderate to very vigorous physical activity; HR: hazard ratio; CI: 95% confidence interval.

**Table S1: Missing information on covariates (N=22,271).
The Tromsø Study 2007-2016.**

Covariate:	Missing, n (%)
Leisure-time physical activity	916 (4)
Exercise frequency	389 (2)
Exercise intensity	1,624 (7)
Exercise duration	1,676 (8)
Waist-height-ratio	172 (1)
Education level	336 (2)
Chronic pain	1,647 (7)
Present and past daily smoking	368 (2)
Average alcohol consumption	390 (2)
Self-reported health	170 (1)

Table S2: Hazard ratios of hand withdrawal on cold-pressor test tolerance according to levels of physical activity by sex^a, using imputed datasets^b. The Tromsø Study 2007-2016.

PA type	HR all (CI)	p, trend	HR women (CI)	HR men (CI)
Leisure-time PA, per unit	0.91 (0.88-0.93)	<0.001	0.93 (0.89-0.96)	0.89 (0.86-0.93)
Sedentary	1		1	1
Light	0.91 (0.86-0.95)		0.95 (0.89-1.02)	0.84 (0.78-0.90)
Moderate	0.83 (0.78-0.89)		0.90 (0.83-0.98)	0.76 (0.70-0.83)
Vigorous	0.70 (0.61-0.80)		0.61 (0.50-0.75)	0.78 (0.66-0.93)
Exercise frequency, per unit	0.98 (0.96-1.01)	0.224	0.96 (0.93-1.00)	1.01 (0.97-1.05)
< 1/week	1		1	1
1-3 times/week	0.99 (0.94-1.04)		0.98 (0.92-1.06)	1.00 (0.93-1.07)
Approximately every day	0.97 (0.91-1.03)		0.93 (0.86-1.01)	1.02 (0.93-1.11)
Exercise intensity, per unit	0.94 (0.91-0.97)	<0.001	0.94 (0.90-0.99)	0.92 (0.88-0.97)
Light	1		1	1
Moderate	0.93 (0.89-0.97)		0.95 (0.90-1.00)	0.90 (0.85-0.96)
Vigorous	0.90 (0.82-1.00)		0.90 (0.78-1.04)	0.92 (0.80-1.06)
Exercise duration, per unit	0.91 (0.88-0.93)	<0.001	0.92 (0.89-0.96)	0.89 (0.85-0.93)
0-29 min.	1		1	1
30-60 min.	0.86 (0.82-0.90)		0.86 (0.81-0.92)	0.87 (0.81-0.93)
>60 min.	0.82 (0.78-0.87)		0.86 (0.80-0.93)	0.79 (0.73-0.86)

^a Cox proportional hazards regression.

^b Multiple imputation with chained equations; predictive mean matching (known nearest neighbours=10), 100 imputed datasets.

Stratified models are adjusted for: sex, age, waist-height-ratio, education, current smoker status, average weekly alcohol consumption, self-reported health and chronic pain. Statistically significant results denoted by **bold**.

Disregarding sex, stratified models use identical adjustments.

PA: physical activity; HR: hazard ratio; CI: 95% confidence interval.

Table S3: Hazard ratios of hand withdrawal on cold-pressor pain tolerance test according to levels of physical activity by chronic pain (yes/no)^a. The Tromsø Study 2007-2016.

PA type	n	Chronic pain ≥ 3 months, yes/no		<i>p</i> ^b
		No	Yes	
		HR (95% CI)	HR (95% CI)	
PA Leisure, per level increase	19,084	0.92 (0.89 - 0.96)	0.90 (0.85 - 0.94)	0.33
Sedentary	2,872	1	1	
Light	11,151	0.88 (0.82 - 0.95)	0.95 (0.88 - 1.04)	0.16
Moderate	4,509	0.86 (0.80 - 0.94)	0.79 (0.71 - 0.88)	0.21
Vigorous	552	0.69 (0.58 - 0.81)	0.78 (0.61 - 1.00)	0.39
Exercise frequency, per level increase	19,388	0.99 (0.95 - 1.23)	0.96 (0.92 - 1.01)	0.45
< 1/wk	3,187	1	1	
1-3 times/wk	11,094	1.00 (0.93 - 1.07)	0.98 (0.90 - 1.06)	0.65
Aprox. every day	5,107	0.98 (0.90 - 1.06)	0.93 (0.85 - 1.03)	0.44
Exercise intensity, per level increase	18,393	0.98 (0.94 - 1.03)	0.91 (0.86 - 0.96)	0.03
Light	7,212	1	1	
Moderate	10,402	0.96 (0.91 - 1.02)	0.92 (0.86 - 0.98)	0.25
Vigorous	779	1.03 (0.91 - 1.16)	0.78 (0.64 - 0.94)	0.02
Exercise duration, per level increase	18,343	0.92 (0.88 - 0.95)	0.89 (0.85 - 0.94)	0.34
0-29 mins	3,681	1	1	
30-60 mins	10,596	0.86 (0.81 - 0.92)	0.85 (0.79 - 0.92)	0.87
>60 mins	4,066	0.84 (0.78 - 0.90)	0.79 (0.72 - 0.88)	0.39
Accelerometry^c:	4922			
Total counts per day		1.00 (0.95 - 1.05)	0.99 (0.92 - 1.05)	0.79
10-minute MVPA minutes		0.97 (0.93 - 1.02)	0.98 (0.92 - 1.06)	0.76

^a Cox proportional hazards regression including two-way interaction terms between chronic pain and physical activity levels.

^b Statistical significance for physical activity*chronic pain interaction term. Significant results in **bold**.

^c Hazard ratios for 1 standard deviation increase.

PA: physical activity; MVPA: moderate to very vigorous physical activity; HR: hazard ratio; CI: confidence interval.

Table S4: Hazard ratios of hand withdrawal on cold-pressor pain tolerance test according to levels of physical activity by moderate-to-severe chronic pain (yes/no)^a. Multiple imputation and complete cases regression. The Tromsø Study 2007-2016.

PA type	ICD11-based ^b moderate-to-severe chronic pain: imputed missing ^c .				ICD11-based moderate-to-severe chronic pain: complete cases.			
	n ^d	No HR (95% CI)	Yes HR (95% CI)	p ^e	n	No HR (95% CI)	Yes HR (95% CI)	p ^e
PA Leisure, per level	17,718	0.87 (0.85 - 0.90)	0.91 (0.86 - 0.97)	0.23	15,563	0.88 (0.85 - 0.92)	0.91 (0.85 - 0.98)	0.46
Sedentary	2,445	1	1	-	2,091	1	1	-
Light	10,273	0.84 (0.79 - 0.90)	0.92 (0.83 - 1.03)	0.09	9,011	0.85 (0.79 - 0.91)	0.93 (0.83 - 1.05)	0.17
Moderate	4,447	0.76 (0.71 - 0.82)	0.83 (0.73 - 0.95)	0.32	3,963	0.78 (0.72 - 0.84)	0.82 (0.71 - 0.96)	0.50
Vigorous	553	0.63 (0.54 - 0.73)	0.78 (0.55 - 1.10)	0.18	498	0.66 (0.56 - 0.77)	0.80 (0.53 - 1.22)	0.39
Exercise frequency, per level	17,718	0.95 (0.92 - 0.99)	0.94 (0.88 - 0.99)	0.67	15,807	0.96 (0.92 - 0.99)	0.95 (0.88 - 1.01)	0.73
< 1/wk	2,693	1	1	-	2,377	1	1	-
1-3 times/wk	10,105	0.95 (0.89 - 1.01)	0.97 (0.88 - 1.08)	0.55	9,059	0.96 (0.89 - 1.02)	0.96 (0.85 - 1.08)	0.95
Aprox. every day	4,920	0.90 (0.84 - 0.97)	0.88 (0.78 - 0.99)	0.74	4,371	0.92 (0.85 - 0.99)	0.90 (0.78 - 1.03)	0.76
Exercise intensity, per level	17,718	0.90 (0.87 - 0.94)	0.87 (0.81 - 0.94)	0.38	15,090	0.93 (0.89 - 0.97)	0.89 (0.82 - 0.97)	0.36
Light	6,842	1	1	-	5,588	1	1	-
Moderate	10,122	0.88 (0.84 - 0.92)	0.89 (0.82 - 0.97)	0.80	8,824	0.90 (0.86 - 0.95)	0.92 (0.83 - 1.01)	0.77
Vigorous	754	0.90 (0.80 - 1.00)	0.67 (0.52 - 0.86)	0.03	678	0.96 (0.85 - 1.08)	0.67 (0.50 - 0.91)	0.03
Exercise duration, per level	17,718	0.90 (0.87 - 0.93)	0.95 (0.89 - 1.01)	0.08	15,155	0.90 (0.83 - 1.16)	0.96 (0.89 - 1.03)	0.12
0-29 mins	3,895	1	1	-	3,046	1	1	-
30-60 mins	9,991	0.87 (0.82 - 0.92)	0.90 (0.82 - 0.99)	0.50	8,689	0.86 (0.81 - 0.91)	0.89 (0.80 - 0.995)	0.56
>60 mins	3,832	0.81 (0.76 - 0.87)	0.91 (0.81 - 1.02)	0.05	3,420	0.81 (0.76 - 0.87)	0.93 (0.81 - 1.07)	0.09
Accelerometry^f:	n/a	-	-	-	5,463			
Total counts per day		-	-	-		0.97 (0.93 - 1.02)	1.03 (0.94 - 1.13)	0.22
10-minute MVPA minutes		-	-	-		0.97 (0.93 - 1.01)	0.97 (0.87 - 1.08)	0.96

^a Cox proportional hazards regression including two-way interaction terms between moderate-to-severe chronic pain and physical activity levels.

^b Moderate-to-severe chronic pain: onset ≥ 3months, intensity >3, impact on ADL >3, bothersomeness >3.

^c Multiple imputation using chained equations with predictive mean matching, number of known nearest neighbours=10.

^d Due to slight sampling variation in imputation, we report group numbers from first imputed dataset here.

^e Statistical significance for physical activity*chronic pain interaction term. Significant results in **bold**.

^f Hazard ratios for 1 standard deviation increase.

ICD: international classification of disease; PA: physical activity; MVPA: moderate to very vigorous physical activity; HR: hazard ratio; CI: confidence interval.